

Recap: ① Some basic definitions of linear model

② Definitions of simple and multiple linear regression.

③ Some concepts of design of experiment.

~~$y = \beta_0 + \beta_1 x_1 + \dots + \beta_p x_p + \epsilon$~~

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_p x_p + \epsilon$$

↑
predictor

β_1, \dots, β_p are called predictor coefficients.

Example:

y_{ij} is the nitrogen concentration of i th plot that received j th level of the fertilizer.

$i = 1, \dots, n$; $j = 1, \dots, a$

$$y_{ij} = \mu + \alpha_j + \epsilon_{ij} \leftarrow \text{error}$$

↑ mean level ↑ effect due to j th level of the fertilizer.

$$\left. \begin{aligned} y_{11} &= \mu + \alpha_1 + \epsilon_{11} \\ \vdots \\ y_{n1} &= \mu + \alpha_1 + \epsilon_{n1} \\ y_{12} &= \mu + \alpha_2 + \epsilon_{12} \\ \vdots \\ y_{n2} &= \mu + \alpha_2 + \epsilon_{n2} \\ \vdots \\ y_{1a} &= \mu + \alpha_a + \epsilon_{1a} \\ \vdots \\ y_{na} &= \mu + \alpha_a + \epsilon_{na} \end{aligned} \right\}$$

$$\begin{pmatrix} y_{11} \\ \vdots \\ y_{na} \end{pmatrix}$$

$$= \begin{bmatrix} 1 & 1 & 0 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & 0 & 0 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & 0 & 0 & \dots & 0 \end{bmatrix}$$

$$\begin{pmatrix} \mu \\ \alpha_1 \\ \alpha_2 \\ \vdots \\ \alpha_a \end{pmatrix} + \begin{pmatrix} \epsilon_{11} \\ \vdots \\ \epsilon_{na} \end{pmatrix}$$

$$\underline{y} = \begin{pmatrix} y_{11} \\ \vdots \\ y_{na} \end{pmatrix}$$

$$\underline{y} = \begin{bmatrix} \underline{1}_n & \underline{1}_n & \dots & \underline{0}_n \\ \vdots & \underline{0}_n & \dots & \vdots \\ \underline{1}_n & \underline{0}_n & \dots & \underline{1}_n \end{bmatrix} \begin{pmatrix} \mu \\ \alpha_1 \\ \vdots \\ \alpha_a \end{pmatrix} + \begin{pmatrix} e_{11} \\ \vdots \\ e_{na} \end{pmatrix}$$

$na \times (a+1)$

$$\underline{1}_n = \underbrace{\begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix}}_{n \text{ times}}, \quad \underline{0}_n = \begin{pmatrix} 0 \\ \vdots \\ 0 \end{pmatrix} \left. \vphantom{\underline{1}_n} \right\} n \text{ times}$$

$$\underline{X} = \begin{bmatrix} \underline{1}_n & \underline{1}_n & \dots & \underline{0}_n \\ \vdots & \underline{0}_n & \dots & \vdots \\ \underline{1}_n & \underline{0}_n & \dots & \underline{1}_n \end{bmatrix}, \quad \underline{\beta} = \begin{pmatrix} \mu \\ \alpha_1 \\ \vdots \\ \alpha_a \end{pmatrix},$$

$$\underline{e} = \begin{pmatrix} e_{11} \\ \vdots \\ e_{na} \end{pmatrix} \Rightarrow \underline{y} = \underline{X} \underline{\beta} + \underline{e}$$

Definition: Kronecker product between two matrices

$\underline{A}_{n \times m}$ and $\underline{B}_{k \times l}$ is ~~given~~ denoted by $\underline{A} \otimes \underline{B}$

and the entries of $\underline{A} \otimes \underline{B}$ is given by the following

$$\left(\begin{array}{cccc} a_{11} \underline{B}_{k \times l} & a_{12} \underline{B}_{k \times l} & \dots & a_{1m} \underline{B}_{k \times l} \\ \vdots & \vdots & \ddots & \vdots \\ a_{n1} \underline{B}_{k \times l} & a_{n2} \underline{B}_{k \times l} & \dots & a_{nm} \underline{B}_{k \times l} \end{array} \right)_{nk \times ml}$$

$$\underline{X} = \left[\begin{array}{cccc} \underline{1}_n & \underline{1}_n & \underline{0}_n & \dots & \underline{0}_n \\ \vdots & \underline{0}_n & \underline{1}_n & & \underline{0}_n \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \underline{1}_n & \underline{0}_n & \underline{0}_n & \dots & \underline{1}_n \end{array} \right] = \left[\underline{I}_a \otimes \underline{1}_n : \underline{I}_a \otimes \underline{1}_n \right]$$

a ~~time~~

$$\underline{I}_a \otimes \underline{1}_n = \begin{pmatrix} 1 \cdot \underline{1}_n & 0 \cdot \underline{1}_n & \dots & 0 \cdot \underline{1}_n \\ 0 \cdot \underline{1}_n & 1 \cdot \underline{1}_n & \dots & 0 \cdot \underline{1}_n \\ \vdots & \vdots & \ddots & \vdots \\ 0 \cdot \underline{1}_n & \dots & \dots & 1 \cdot \underline{1}_n \end{pmatrix}$$

$$\underline{1}_a \otimes \underline{1}_n = \begin{pmatrix} 1 \cdot \underline{1}_n \\ 1 \cdot \underline{1}_n \\ \vdots \\ 1 \cdot \underline{1}_n \end{pmatrix}$$

$$\underline{X} = \left[\underline{1}_a \otimes \underline{1}_n : \underline{I}_a \otimes \underline{1}_n \right]$$

$$\underline{y} = \underline{X} \underline{\beta} + \underline{e} \quad (\text{One way Anova model})$$

Example: There are two different fertilizers and the first one can be applied at a different levels, the second one can be applied at b different levels. Each combination of fertilizers is applied to n different agricultural plots.

$$y_{ijk} = \mu + \alpha_i + \beta_j + e_{ijk}$$

y_{ijk} = Nitrogen conc. of the kth plot that received i-th level of fertilizer 1 and jth level of fertilizer 2.

$$k = 1, \dots, n.$$

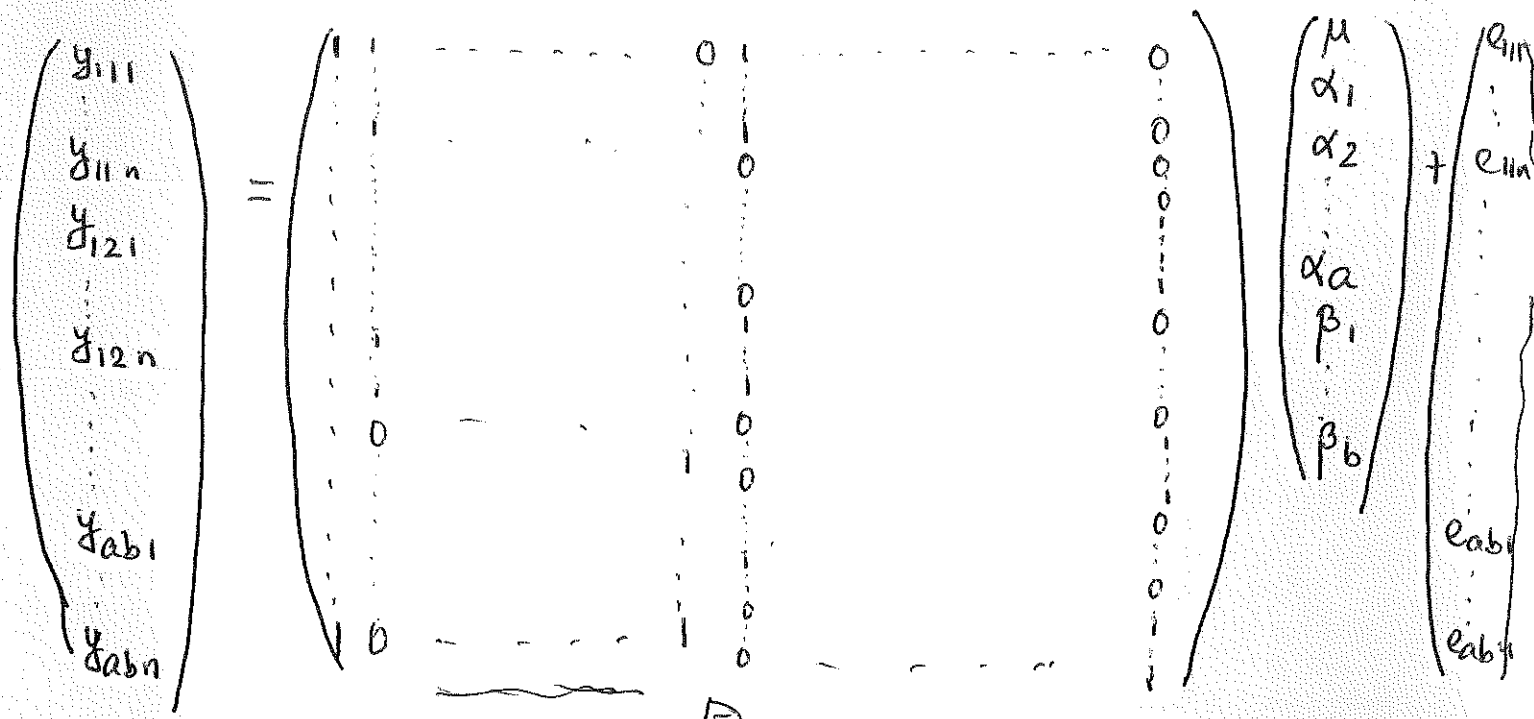
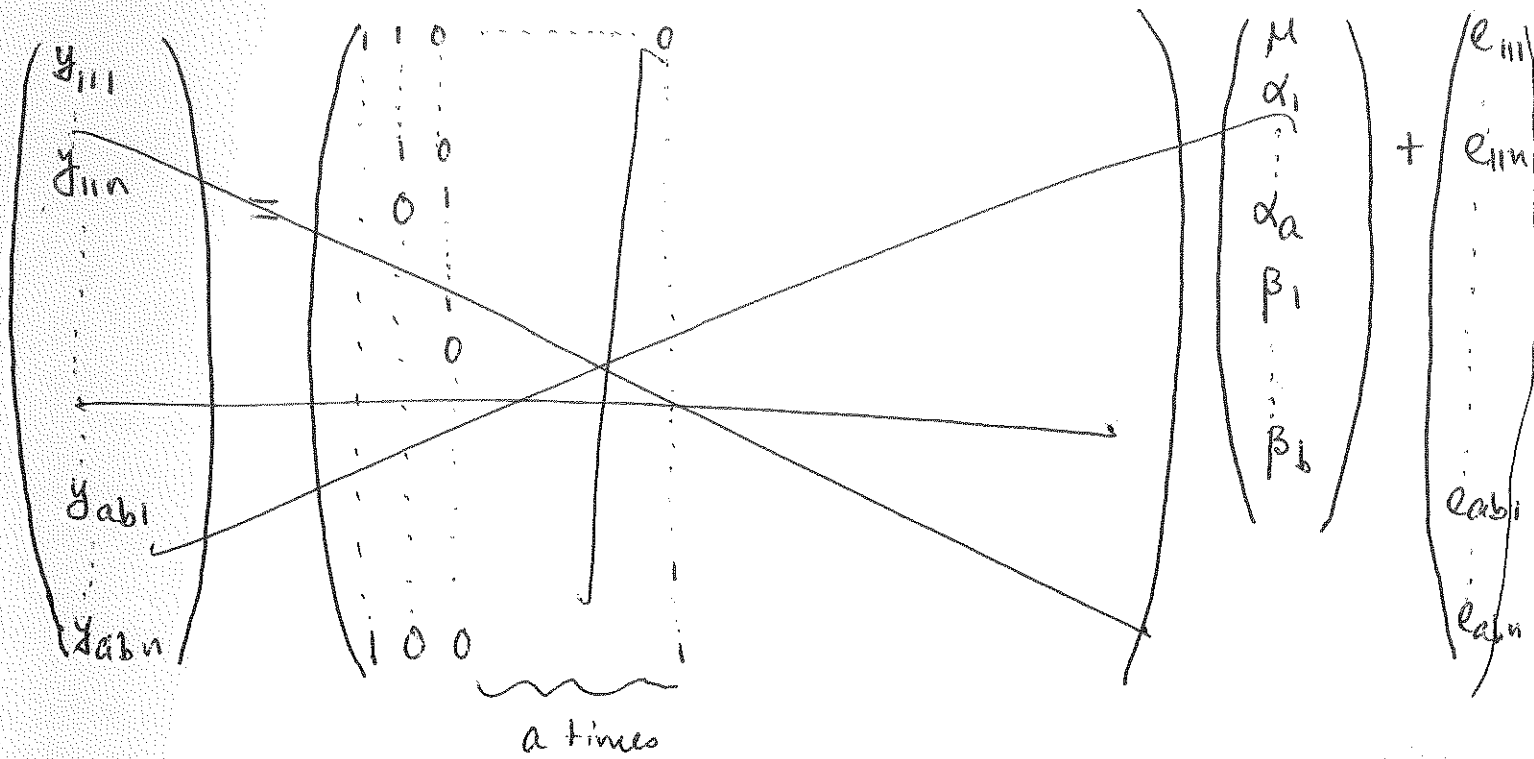
$$y_{111} = \mu + \alpha_1 + \beta_1 + e_{111}$$



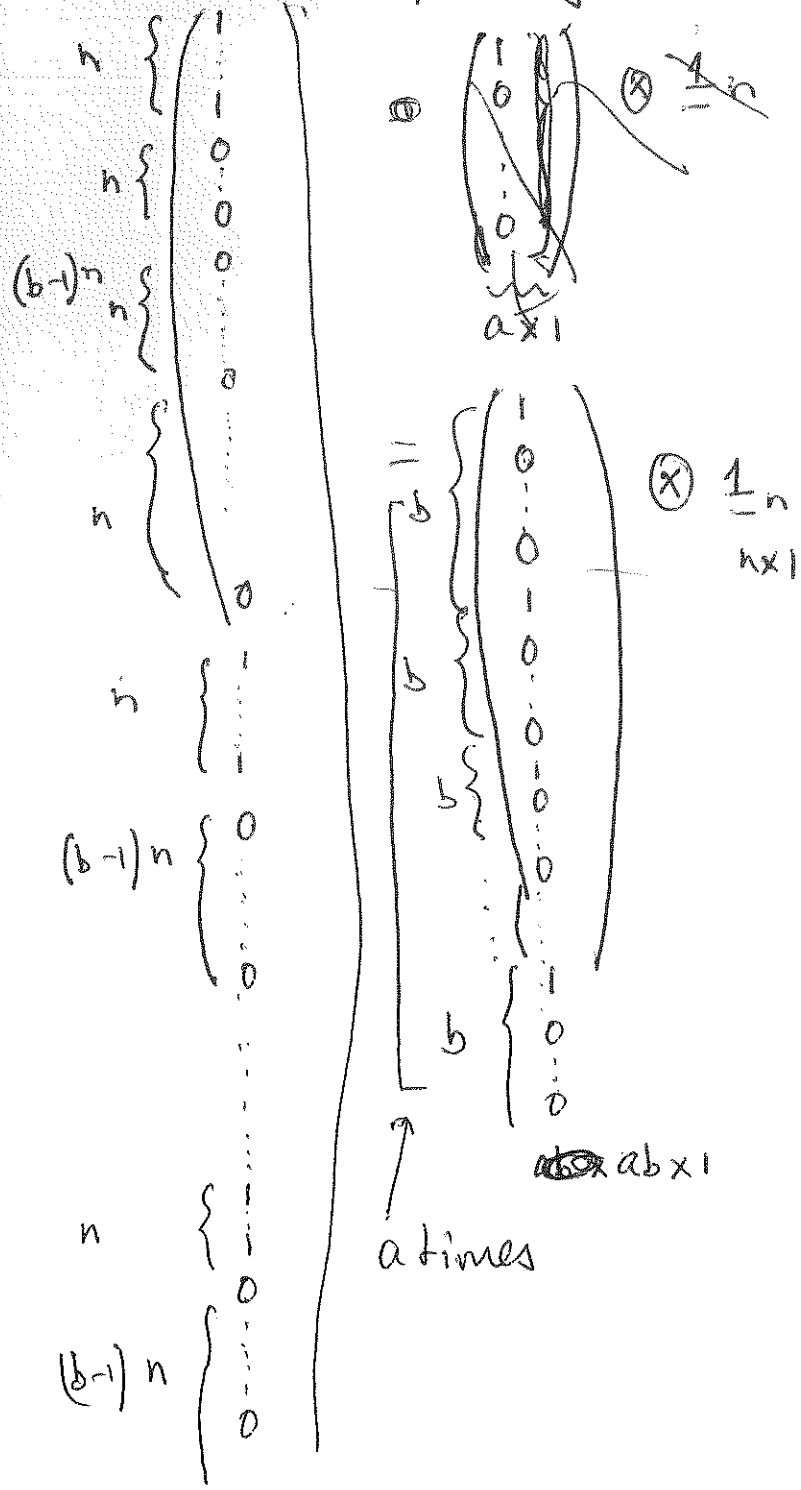
$$y_{11n} = \mu + \alpha_1 + \beta_1 + e_{11n}$$

$$y_{a11} = \mu + \alpha_a + \beta_1 + e_{a11}$$

$$y_{abn} = \mu + \alpha_a + \beta_b + e_{abn}$$



Column corresponding to β_1



$$y_{111} = \mu + \alpha_1 + \beta_1 + e_{111}$$

$$\vdots$$

$$y_{11n} = \mu + \alpha_1 + \beta_1 + e_{11n}$$

$$y_{121} = \mu + \alpha_1 + \beta_2 + e_{121}$$

\vdots

$$y_{12n} = \mu + \alpha_1 + \beta_2 + e_{12n}$$

\vdots

$$y_{1b1} = \mu + \alpha_1 + \beta_b + e_{1b1}$$

\vdots

$$y_{1bn} = \mu + \alpha_1 + \beta_b + e_{1bn}$$

$$y_{211} = \mu + \alpha_2 + \beta_1 + e_{211}$$

\vdots

$$y_{21n} = \mu + \alpha_2 + \beta_1 + e_{21n}$$

\vdots

\vdots

\vdots

\vdots

$$y_{abi} = \mu + \alpha_a + \beta_b + e_{abi}$$

\vdots

\vdots

$$y_{abn} = \mu + \alpha_a + \beta_b + e_{abn}$$

$$\underline{X} = \begin{bmatrix} \underline{1}_n & \underline{1}_n & \underline{0}_n & \dots & \underline{0}_n & \underline{1}_n & \dots & \underline{0}_n \\ \underline{1}_n & \dots & \dots & \dots & \dots & \underline{0}_n & \dots & \dots \\ \dots & \underline{1}_n & \underline{0}_n & \dots & \dots & \underline{0}_n & \dots & \dots \\ \dots & \underline{0}_n & \underline{1}_n & \dots & \dots & \underline{1}_n & \dots & \dots \\ \underline{1}_n & \underline{0}_n & \underline{1}_n & \dots & \dots & \underline{0}_n & \dots & \dots \\ \underline{1}_n & \dots & \underline{0}_n & \dots & \dots & \underline{0}_n & \dots & \dots \\ \dots & \dots & \dots & \dots & \underline{0}_n & \dots & \dots & \dots \\ \dots & \underline{0}_n & \underline{0}_n & \dots & \underline{1}_n & \underline{1}_n & \dots & \underline{0}_n \\ \underline{1}_n & \underline{0}_n & \underline{0}_n & \dots & \underline{1}_n & \underline{0}_n & \dots & \underline{0}_n \\ & & & & & & & \underline{1}_n \end{bmatrix}$$

$$= \left[\underline{1}_{ab} \otimes \underline{1}_n : \underline{I}_a \otimes \underline{1}_{nb} : \begin{pmatrix} \underline{I}_b \otimes \underline{1}_n \\ \underline{I}_b \otimes \underline{1}_n \\ \vdots \\ \underline{I}_b \otimes \underline{1}_n \end{pmatrix} \right]$$

$$= \left[\underline{1}_{ab} \otimes \underline{1}_n : \underline{I}_a \otimes \underline{1}_{nb} : \underline{1}_a \otimes \underline{I}_b \otimes \underline{1}_n \right]$$

$$\underline{y} = \underline{X} \underline{\beta} + \underline{e} \quad \underline{\beta} = \begin{pmatrix} \mu \\ \alpha_1 \\ \vdots \\ \alpha_a \\ \beta_1 \\ \vdots \\ \beta_b \end{pmatrix}$$

How to estimate the coefficient vector $\underline{\beta}$

$$\underline{y} = X\underline{\beta} + \underline{e}$$

Least square estimator

Def: Let \underline{h} be a vector of dimension n .

$$\|\underline{h}\| = \sqrt{\underline{h}^T \underline{h}}$$

$$\underline{h} = (h_1, \dots, h_n)^T$$

$$= \sqrt{\sum_{i=1}^n h_i^2}$$

Distance between two vectors \underline{x} , \underline{y} both of dimension $n \times 1$ is given by

$$\|\underline{x} - \underline{y}\| = \sqrt{(\underline{x} - \underline{y})^T (\underline{x} - \underline{y})}$$

$$\underline{x} = \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix}, \quad \underline{y} = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}$$

$$\text{then } \|\underline{x} - \underline{y}\| = \sqrt{\sum_{i=1}^n (x_i - y_i)^2}$$

Least square estimator of $\underline{\beta}$ is obtained by minimizing

$$Q(\underline{\beta}) = \|\underline{y} - X\underline{\beta}\|^2 = (\underline{y} - X\underline{\beta})^T (\underline{y} - X\underline{\beta})$$

w.r.t. $\underline{\beta}$.

$$\text{write } \frac{\partial Q(\underline{\beta})}{\partial \underline{\beta}} = \underline{0}$$

Derivative w.r.t. a vector

Let $\underline{a}, \underline{b}$ be $p \times 1$ vectors and \underline{A} be a $p \times p$ matrix. Then,

$$(1) \quad \frac{\partial (\underline{b}^T \underline{a})}{\partial \underline{b}} = \frac{\partial (\underline{a}^T \underline{b})}{\partial \underline{b}} = \underline{a}$$

$$(2) \quad \frac{\partial (\underline{b}^T \underline{A} \underline{b})}{\partial \underline{b}} = (\underline{A} + \underline{A}^T) \underline{b}$$

$$\begin{aligned} Q(\underline{\beta}) &= (\underline{y} - \underline{X} \underline{\beta})^T (\underline{y} - \underline{X} \underline{\beta}) \\ &= \underline{y}^T \underline{y} - \underline{y}^T \underline{X} \underline{\beta} - \underline{\beta}^T \underline{X}^T \underline{y} + \underline{\beta}^T \underline{X}^T \underline{X} \underline{\beta} \end{aligned}$$

$$\begin{aligned} \frac{\partial Q(\underline{\beta})}{\partial \underline{\beta}} &= \underline{0} - \underline{X}^T \underline{y} - \underline{X}^T \underline{y} + (\underline{X}^T \underline{X} + (\underline{X}^T \underline{X})^T) \underline{\beta} \\ &= -2 \underline{X}^T \underline{y} + 2 \underline{X}^T \underline{X} \underline{\beta} \end{aligned}$$

$$\frac{\partial Q(\underline{\beta})}{\partial \underline{\beta}} = \underline{0} \iff \underline{X}^T \underline{X} \underline{\beta} = \underline{X}^T \underline{y} \quad (\text{Normal Equations})$$

Let \underline{X} be of dimension $n \times p$, \underline{y} ~~be~~ be of dimension $n \times 1$, $\underline{\beta}$ is of dimension $p \times 1$

$$\Rightarrow \begin{matrix} \underline{X}^T & \underline{X} & \underline{\beta} & = & \underline{X}^T \underline{y} \\ p \times n & n \times p & p \times 1 & & p \times n \quad n \times 1 \end{matrix}$$

When $(\underline{X}^T \underline{X})$ can be inverted these equations has a unique solution $\underline{\beta} = (\underline{X}^T \underline{X})^{-1} \underline{X}^T \underline{y}$

But when $\underline{X}^T \underline{X}$ is not invertible, then these

equations have infinitely many solutions.

Example: $y_i = \beta_0 + \beta_1 x_i + e_i$, $i = 1, \dots, n$

$$\underline{y} = \underline{X} \underline{\beta} + \underline{e} \quad \underline{X} = \begin{pmatrix} 1 & x_1 \\ \vdots & \vdots \\ 1 & x_n \end{pmatrix}, \quad \underline{y} = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}$$

$$\underline{e} = \begin{pmatrix} e_1 \\ \vdots \\ e_n \end{pmatrix}, \quad \underline{\beta} = \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix}$$

$$\underline{X}^T \underline{X} \underline{\beta} = \underline{X}^T \underline{y}$$

$$\underline{X}^T \underline{X} = \begin{pmatrix} 1 & \dots & \dots & 1 \\ x_1 & \dots & \dots & x_n \end{pmatrix} \begin{pmatrix} 1 & x_1 \\ \vdots & \vdots \\ \vdots & \vdots \\ 1 & x_n \end{pmatrix} = \begin{pmatrix} n & \sum_{i=1}^n x_i \\ \sum_{i=1}^n x_i & \sum_{i=1}^n x_i^2 \end{pmatrix}$$

we will see later that $(\underline{X}^T \underline{X})^{-1}$ exists if and only if not all x_i 's are the same.

Thus the Normal ~~Eq~~ Equations have a unique solution if not all x_i 's are equal.